



Detecting Resource Limitation in a Large Herbivore Population Is Enhanced With Measures of Nutritional Condition

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Resource limitation at the population level is a function of forage quality and its abundance relative to its per capita availability, which in turn, determines nutritional condition of individuals. Effects of resource limitation on population dynamics in ungulates often occur through predictable and sequential changes in vital rates, which can enable assessments of how resource limitation influences population growth. We tested theoretical predictions of bottom-up (i.e., resource limitation) forcing on moose (*Alces alces*) through the lens of vital rates by quantifying the relative influence of intrinsic measures of nutritional condition and extrinsic measures of remotely sensed environmental data on demographic rates. We measured rates of pregnancy, parturition, juvenile, and adult survival for 82 adult females in a population where predators largely were absent. Life stage simulation analyses (LSAs) indicated that interannual fluctuations in adult survival contributed to most of the variability in λ . We then extended the LSA to estimate vital rates as a function of bottom-up covariates to evaluate their influence on λ . We detected weak signatures of effects from environmental covariates that were remotely sensed and spatially explicit to each seasonal range. Instead, nutritional condition strongly influenced rates of pregnancy, parturition, and overwinter survival of adults, clearly implicating resource limitation on λ . Our findings depart from the classic life-history paradigm of population dynamics in ungulates in that adult survival was highly variable and generated most of the variability in population growth rates. At the surface, lack of variation explained by environmental covariates may suggest weak evidence of resource limitation in the population, when nutritional condition actually underpinned most demographics. We suggest that variability in vital rates and effects of resource limitation may depend on context more than previously appreciated, and density dependence can obfuscate the relationships between remotely sensed data and demographic rates.

Keywords: bottom-up, life stage simulation analysis, survival, ungulate, density dependence, vital rate, moose, population

INTRODUCTION

Empirical evidence along with perceived declines in abundance of moose (*Alces alces*) across the southern distribution of their range has resulted in heightened interest in identifying factors limiting population growth. Recent studies indicate that nutritional limitation stemming from suboptimal habitat conditions, parasite abundance, and thermal stress associated with climatic warming and drying are contributing to poor population performance in the Intermountain West (Becker et al., 2010; Monteith et al., 2015), Minnesota (Murray et al., 2006; Lenarz et al., 2009; DelGiudice et al., 2011), Michigan (Peterson, 1999; Dodge et al., 2004), and in the northeastern United States (Musante et al., 2010; Bergeron et al., 2013). Confounding the evidence for nutritional limitation is the recolonization of large carnivores, namely, grizzly bears (*Ursus arctos horribilis*) and gray wolves (*Canis lupus*), which, in the Intermountain West, occurred concurrently with changes in temperature and precipitation patterns (Bangs and Fritts, 1996; Berger et al., 2001). The relative importance of predation and resource limitation (i.e., forage quality and abundance, its per capita availability) in structuring dynamics of animal populations has been debated for decades, which was catalyzed partially by the Green World Hypothesis (GWH; Hairston et al., 1960). For prey populations, the continuum of top-down and bottom-up forcing implies that trade-offs occur between acquisition of resources for survival and susceptibility to predation (McNamara and Houston, 1987), mediated by proximity to carrying capacity (K) (Sinclair and Krebs, 2002; Pierce et al., 2012; Bowyer et al., 2013). Nonetheless, declines in recruitment of young and abundance of moose are also apparent in regions where large carnivores are absent, further suggesting a predominant role of nutritional limitation in populations near the southern distribution of moose. The life-history paradigm (Eberhardt, 2002) exhibited by ungulates offers a useful predictive framework to assess the influence of resource limitation on animal populations (Gaillard et al., 2000; Eberhardt, 2002; Bonenfant et al., 2009; Monteith et al., 2014). Understanding how nutritional limitation is expressed in populations near K can enhance predictive models that aim to determine the relative influence of top-down and bottom-up forces on ungulate population dynamics.

The proximity of prey populations to K has considerable influence on the relative strength of top-down and bottom-up forces that structure ungulate demography (Kie et al., 2003; Bonenfant et al., 2009). For example, populations near K tend to be in poor nutritional condition with reduced fecundity and survival (McCullough, 1979; Simard et al., 2008). Nevertheless, predator populations can reduce density-dependent feedbacks by killing prey and reducing intraspecific competition (Boyce et al., 1999). Consequently, we expect vital rates and the factors that underpin them to differ between populations regulated by top-down and bottom-up forcing (Pierce et al., 2012). For example, in prey populations under strong top-down forcing, intraspecific competition should be reduced and nutritional condition should increase (Bowyer et al., 2014). In this instance, predation should dampen signals of resource limitation by pushing prey populations below K (Owen-Smith and Mills, 2006;

Hopcraft et al., 2010). In contrast, most populations without predators should exhibit strong signals of resource limitation as density increases, and vital rates should respond accordingly (Coulson et al., 2001; Stewart et al., 2005; Monteith et al., 2014). Indeed, interactions between resource limitation and predation can obscure their relative influence at the population level. Thus, a hypothesis-driven approach to assessing variability in vital rates in populations exposed to few predators can help tease apart and detect the mechanisms of resource limitation, advancing our understanding of the influence of predators on ungulate prey (Gaillard et al., 2000; Bowyer et al., 2013).

Resource limitation on population dynamics in ungulates often occurs through predictable and sequential changes in vital rates: reduced survival of young followed by increased age at first reproduction, reduced rates of pregnancy, parturition, and reduced survival of prime-aged adults (i.e., the “Eberhardt Model”; Eberhardt, 2002). Typically, adult survival is high and largely invariant, and the greater variability in juvenile survival commonly drives population dynamics (Gaillard et al., 1998, 2000; Gaillard and Yoccoz, 2003; Raithel et al., 2007). With a reduced energetic buffer against the vagaries of weather, populations exposed to resource limitation may be more sensitive to weather (Portier et al., 1998; Kie et al., 2003; Hansen et al., 2019), and consequently, environmental variability can lead to dramatic changes in population growth rate (λ) for resource-limited populations (Coulson et al., 2001). In contrast, reduced intraspecific competition in response to top-down forcing can weaken environmental signatures on vital rates, because animals possess an energetic buffer against them (Bowyer et al., 2014).

Assessing nutritional status is critical for understanding the influence of habitat alterations on populations, because the nutritional condition of an individual is an integrated measure of energetic gains and losses relative to food quality and abundance (Cook et al., 2007; Monteith et al., 2014). Nutritional condition forms the foundation for life history of individuals and thus affects nearly every demographic component of a population leading to net effects on population growth (Monteith et al., 2014). Remotely sensed data [e.g., normalized difference vegetation index (NDVI) and standardized precipitation index] are increasingly being used as a proxy for assessing resource limitation in ungulate populations (Pettorelli et al., 2011; Monteith et al., 2015), and their accessibility, broad spatial and temporal availability makes them an attractive tool. Nevertheless, remotely sensed data do not discriminate between the species and quality of forage selected by the animal, and per capita availability is particularly difficult to quantify. Thus, linking relationships between resource limitation and remotely sensed data can be noisy or weak. Measures of nutritional condition inherently integrate density dependence because of its effect on per capita availability of food, which is realized in the nutritional status of an animal and not possible to measure through simple measures of habitat. Moreover, directly quantifying habitat quality is exceedingly difficult and often cost prohibitive (DeYoung et al., 2000; Stephenson et al., 2006). An integrated measure of habitat quality and food availability is possible by measuring nutritional status because animal condition is a direct product of its environment (Franzmann, 1977; Parker et al., 2009).

Predictive equations for estimating ingesta-free body fat (IFBFat) of moose based on ultrasonography measurements of maximum depth of rump fat are quite accurate ($r^2 = 0.96$; Stephenson et al., 1998), and thus represent a meaningful measurement of nutritional condition.

We sought to quantify how resource limitation influences population dynamics through the lens of variability in vital rates of moose in the Sublette Herd, located in a mountainous region of western Wyoming, United States. Further, we wanted to understand the relative context in which measures of nutritional condition would compare with remotely sensed environmental covariates in predicting vital rate responses. We estimated demographic rates of individuals to evaluate predictions of variability in vital rates and their relative influence on λ (Gaillard et al., 1998) in the context of Eberhardt's model (Eberhardt, 2002). We did not incorporate population size into demographic models, because annual minimum counts and juvenile ratios were stable before, during, and after the study period (Wyoming Game and Fish Department, unpublished data). Predators were largely absent and female harvest was low; therefore, we expected negative effects of density dependence to be operating and moose to be sensitive to environmental signals of resource limitation (i.e., summer drought and NDVI). Specifically, we expected (1) reductions in survival of young, and fecundity of adults if resource limitation was occurring (Eberhardt, 2002) and (2) invariant survival of adults, relative to other vital rates (Gaillard et al., 1998). Further, we used individual measures of nutritional condition and remotely sensed data to quantify the ability to detect resource limitation, should it occur. We then used deterministic matrix models and life stage simulation analysis (LSA; Wisdom et al., 2000) to assess vital rate contributions to population-level responses to metrics of resource limitation. Additionally, our study provides insight into the interplay among bottom-up forces as they manifest through the life-history characteristics of moose in a variable environment.

MATERIALS AND METHODS

Study Area

We monitored 82 adult female moose in the Sublette Herd from February 2011 until August 2014 in the Upper Green River basin and eastern foothills of the Wyoming Range (42.8653 degrees North, 110.0708 degrees West; **Figure 1**). Land ownership was 60% private and 40% public lands. During winter, moose occupied riparian areas at low to mid elevations (1866–2150 m), consisting mainly of willow (*Salix* spp.) interspersed with cottonwood (*Populus* spp.), subalpine fir (*Abies lasiocarpa*), Douglas fir (*Pseudotsuga menziesii*), and aspen (*Populus tremuloides*). Homogenous and mixed forests of subalpine fir, lodgepole pine (*Pinus contorta*), aspen, Engelmann spruce (*Picea engelmannii*), and Douglas fir also occurred throughout higher elevations along the foothills of the Wyoming Range. Winter ranges of moose typically spanned the relatively flat floodplains of the Green River, as well as Cottonwood, Horse, and Beaver Creeks, which were dominated by willow and intermixed with cottonwoods. Approximately half of the moose

were resident with overlapping summer and winter seasonal ranges, while migratory individuals traveled short distances (5–20 miles) to higher elevations or other tributaries of the Green and Hoback Rivers containing stands of willow, subalpine fir, lodgepole pine, and Douglas fir interspersed with stands of aspen, limber pine (*Pinus flexilis*) and whitebark pine (*Pinus albicaulis*). Residents occupied willow communities, aspen forests, or mixed-conifer and aspen forest throughout the year. The climate is characterized by short, dry summers and long, cold winters.

The study area was host to four wild ungulates, cougars (*Puma concolor*), and American black bears (*Ursus americanus*) (Buskirk, 2016). Elk (*Cervus canadensis*) were the most numerous ungulate. Mule deer (*Odocoileus hemionus*) and pronghorn (*Antilocapra americana*) occurred throughout. Gray wolves (*C. lupus*) and grizzly bears (*U. arctos*) were rarely documented in the study area.

Captures

We captured moose via helicopter net-gunning (Native Range Capture Services, Inc.) without immobilization agents, and blindfolded, hobbled, and restrained in a sternal-recumbent position. We fit 72 moose with GPS store-on-board collars (TGW-3700 and -4700, Telonics, Inc., Mesa, AZ, United States), and fit 12 with GPS satellite-uplink collars (various D-cell models, North Star Science and Technology, LLC, King George, VA, United States). Telonics collars recorded hourly locations, and North Star collars recorded locations every 3–5 h, depending on the model. Both Telonics and North Star collars were programmed to release from the animal approximately 2 years after deployment. For 13 moose in 2013 with expiring GPS collars, we deployed VHF-collars (M2230B, Advanced Telemetry Systems, Isanti, MN, United States) to continue demographic monitoring. One experienced investigator (K. L. Monteith) collected all data on nutritional condition from 2012 to 2014 ($n = 70$ individuals; nutritional condition estimates were not available for 2011). We first assessed condition of animals by palpation and acquired a modified rBCS, which was based on the extent (cm) to which the index finger could be inserted on the caudal side of the sacro-sciatic ligament. This rBCS was analogous to that validated for elk (Cook et al., 2001) and mule deer (Cook et al., 2007), which were highly correlated with percent IFBFat ($r^2 > 0.88$). We measured depth of rump fat and thickness of the bicep femoris using the electronic calipers (± 0.1 cm) of a Bantam II portable ultrasound device (E.I. Medical Imaging, Loveland, CO, United States) with a 5-MHz linear-array transducer. We measured maximum depth of subcutaneous rump fat (MAXFat) along a line parallel to the spine and cranial to the ischial tuber (pin bone), which occurred immediately adjacent to the cranial process of the ischial tuber (Stephenson et al., 1998). We used subcutaneous rump fat to estimate percent IFBFat for moose with measurable fat. For animals without subcutaneous fat, we used rBCS to estimate percent IFBFat using the linear relationship between IFBFat and rBCS of moose with measurable rump fat (Cook et al., 2010; Jesmer et al., 2017). During captures for years 2013–2014, we removed the right-incisiform canine to estimate age of each moose (Boertje et al., 2015). For study animals that died before

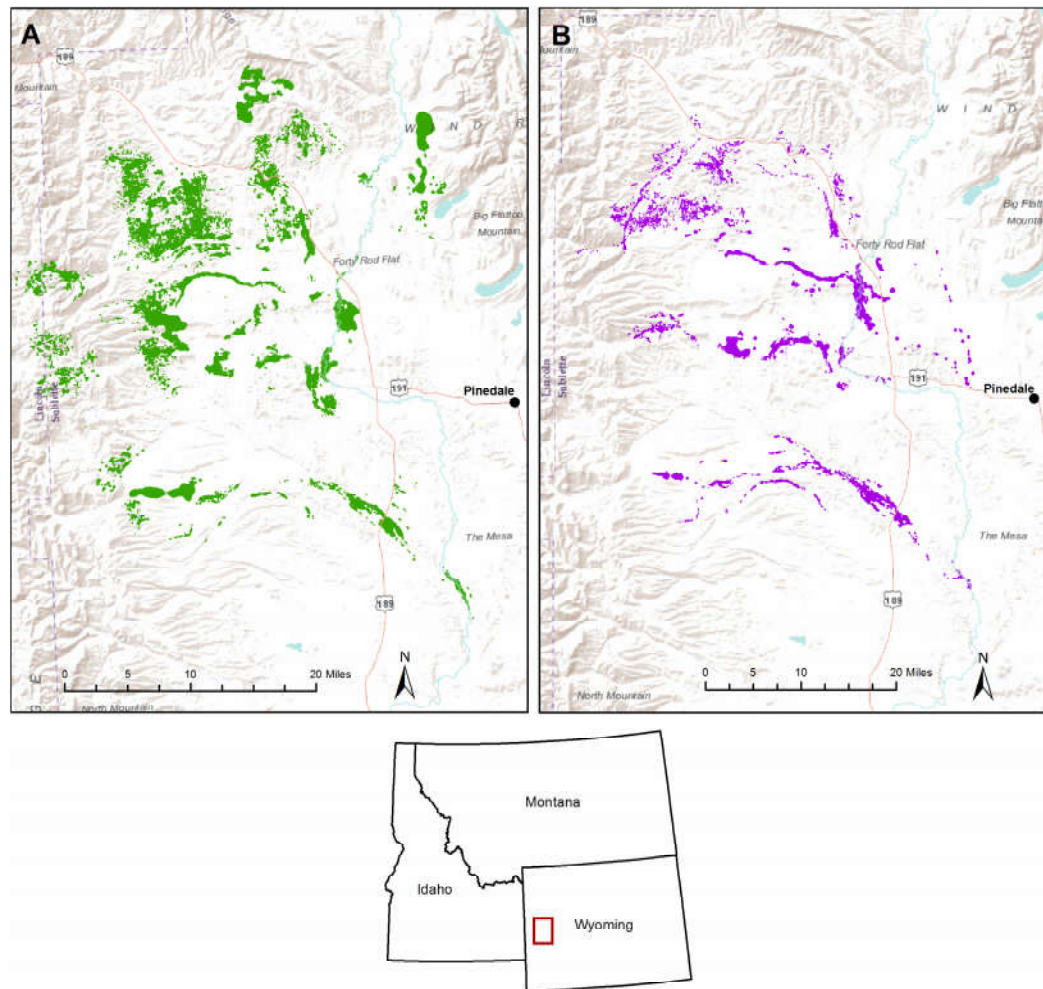


FIGURE 1 | Map of study area showing a sample ($n = 63$) of seasonal home ranges of adult (>2.5 years), GPS-collared moose from 2012 and 2013, derived from dynamic Brownian bridge movement models for summer (A) and winter (B) in Sublette County of western Wyoming, United States.

2013, we collected the I-1 incisor from the mortality site. All age estimates were conducted via cementum annuli by Matson's Laboratory, Milltown, MT, United States. We omitted yearlings from all analyses (<2 and ≥ 1 years), because their vital rates can be lower or more variable compared with reproductive adults (Gaillard et al., 2000; Bonenfant et al., 2009). We collected 20 ml of blood from each female by jugular venipuncture. For analyses, we omitted all mortalities that occurred within 2 weeks of capture ($n = 4$). The University of Wyoming Institutional Animal Care and Use Committee approved procedures for captures and handling (protocol number 20140124JG00057).

Seasonal Ranges of Moose

We delineated summer and winter ranges of each moose using characteristics of net squared displacement. First, we identified moose as migratory if they exhibited directional movement between distinct, non-overlapping winter and summer ranges. Second, we visually inspected plots of net squared displacement of migratory animals to identify start and end dates of migration.

Migratory moose typically left winter ranges between late April and late May, and arrived on summer ranges between early and late June, returning to winter ranges around mid to late December. If seasonal ranges overlapped between winter and summer, we classified moose as residents and used the median start and end migration dates from migratory moose relative to year to define the temporal extent of winter and summer ranges. Net squared displacement calculations were derived from GPS-collar locations by calculating the squared distance between the first location identified from capture and every subsequent location along the annual travel path (Bunnefeld et al., 2011). After identifying migration dates, we delineated seasonal range polygons for moose with GPS collars using 95% contours derived from dynamic Brownian bridge movement models (dBBMMs; Kranstauber et al., 2012; see **Figure 1** for examples). We specified all dBBMMs with a location error of 20 m, raster cell size of 20 m, window size of 31 locations, and a margin of 11 locations. Fix rates of the GPS collars varied, depending on the year and season. For each of the 13 moose in our study that wore a VHF collar after

wearing a GPS collar for two or more years, we used the extent of the seasonal range from the most recent year of the GPS collar to represent seasonal ranges during years the individual was fitted with a VHF collar.

Monitoring of Vital Rates

Pregnancy and Parturition

We separated red blood cells from serum collected at capture via centrifugation and submitted the latter to BioTracking LLC to determine pregnancy status using the presence of Pregnancy-Specific Protein B (PSPB; Huang et al., 2000). We used a PSPB cut-off value of ≥ 2.5 ng/ml to identify pregnant females with viable fetuses for each sample (Josh Branen, BioTracking LLC, personal communication). To identify parturition events, we evaluated females deemed pregnant during capture, observing collared moose aerially with a Bell-47 helicopter between 9 and 13 June, shortly after peak parturition (Schwartz, 1998, pp. 141–171; Poole et al., 2007). We calculated parturition rates as the proportion of pregnant individuals that were parturient. We calculated twinning rates as the proportion of parturient females observed with twins during the June survey.

Survival of Young

Ungulates are most vulnerable to mortality during the first 6 weeks of life (i.e., the neonatal phase), after which mortality typically decreases until winter (Ballard et al., 1991; Barber-Meyer et al., 2008). Thus, we distinguished between survival of neonates (parturition to mid-July) and survival of juveniles (mid-July to February). Approximately 4–8 weeks after parturition surveys, we relocated adult females from a helicopter to record presence and estimate survival of neonates at heel. The following February we estimated survival of juveniles by relocating females that had at least one neonate at heel in July from the air. We used Kaplan–Meier estimators (Pollock et al., 1989) to calculate annual survival rates of neonates and juveniles.

Adult Survival

We monitored the survival of collared moose with monthly fixed-wing flights and recovered collars from the field when logistically feasible. For GPS-collared moose, we verified date of death by visually inspecting fix locations relative to mortality site or dropped collar. We calculated annual survival rates at monthly intervals using a Kaplan–Meier estimator, with the biological year starting 1 June and ending 31 May.

Covariates

To evaluate the influence of metrics of resource limitation on each vital rate, we used IFBBat measurements and extracted a suite of covariates annually from seasonal ranges of each moose (Table 1). For newly captured moose (i.e., individuals for which the previous summer and winter ranges were not known), we used the seasonal range from the current year to represent the seasonal range for the previous year, because moose in the Greater Yellowstone Ecosystem are faithful to their seasonal ranges (Vartanian, 2011). For all environmental covariates, we used the spatially weighted mean of the raster value extracted from each seasonal range (i.e., polygon derived from

the 95% isopleth of dBBMM from GPS-collared moose). Some moose died during the same winter they were first captured, and therefore, had no measurable summer range. To avoid bias associated with censoring these individuals because we lacked data for their summer range, we applied the annual mean value of each environmental covariate extracted from all other summer home ranges. We conducted all data extraction and modeling with R version 4.0.2 (R Core Team, 2020).

Resource Limitation

Length of spring and growing season can affect fecundity and survival in ungulates (Herfindal et al., 2006; Hamel et al., 2010; Table 1). As proxies for environmental effects and characteristics of summer ranges for moose, we used NDVI from the MODIS terra satellite (8-day temporal resolution) to calculate the annual length of spring (start to end of spring-green up of vegetation) and growing season (start of spring-green up until senescence of vegetation) in days (Table 1). We calculated NDVI annually from 2011 to 2014 using the MOD09Q1 data product from the MODIS terra satellite and followed the same protocol for processing NDVI as outlined in Merkle et al. (2016). We extracted length of spring and growing season in days from each summer range using the spatially-weighted mean of all cells overlapping the 95% dBBMM isopleth of the home range. We defined length of spring as the number of days from the start to end of vegetation green-up and length of the growing season from the start of vegetation green-up until the start of fall senescence of vegetation. Length of spring usually included the months of May and June, whereas the length of the growing season usually included months May to September. To evaluate responses of vital rates to NDVI metrics, we used either the value from current year, in addition to or in place of the previous year's value, depending on the vital rate. For example, pregnancy status was measured in February, so we used the value from the previous year to account for carry-over effects (Monteith et al., 2014, 2015). For survival of neonates (June–July), we evaluated length of spring from the current year and evaluated length of the growing season (calculated at the end of summer) from the previous year. We used the value from the previous year for length of spring and length of growing season to evaluate survival of juveniles (July–February). Although fires have occurred in Sublette, we did not assess their influence on vital rates because proportional overlap with home ranges was low.

We estimated drought using PRISM raster data of the Palmer Drought Severity Index (PDSI; Palmer, 1965; Table 1) extracted from the seasonal home range of each individual. Warm temperatures during late winter associated with reduced snow cover have been correlated with winter tick (*Dermacentor albipictus*) abundance the subsequent year (DelGiudice et al., 1997), which can decrease body condition of moose, especially juveniles, through loss of blood and hair (Samuel, 2007).

We estimated seasonal drought annually from the PDSI (Palmer, 1965) using rasters of 4-km² resolution from PRISM (Daly et al., 1994). The annual PDSI per summer home range of moose was calculated at the end of summer; therefore, we applied the PDSI value from the previous year to evaluate the influence of drought on pregnancy (February), parturition

TABLE 1 | Descriptions of covariates evaluated for their relative influence on probability of pregnancy, parturition, survival of juveniles (July–February), and overwinter adults (January–May) of Sublette moose from 2012 to 2014 moose in western Wyoming, United States.

Covariate (data source)	Spatial resolution (temporal resolution)	Influence	Description	Home range tested	Vital rates tested	Justification
Spring length (MODIS, NDVI ^a)	250 m (8-day)	Forage quality and digestibility	Length of spring in days, from the start to end of spring green-up of vegetation from previous or current year	Summer	All	Pettorelli et al., 2007; Hebblewhite et al., 2008; Hamel et al., 2009b; Monteith et al., 2015
Growing season length (MODIS, NDVI ^a)	250 m (8-day)	Forage quality and digestibility	Length of the growing season in days, from start of spring green-up to start of fall senescence of vegetation from previous or current year	Summer	All	Hjeljord and Histol, 1999; Ericsson et al., 2002; Herfindal et al., 2006
Summer Palmer Drought Severity Index (PRISM ^b)	4 km (Monthly)	Forage quality and digestibility	Home range value of Palmer Drought Severity Index experienced across all summer ranges from 1 June to 31 August from previous or current year	Summer	All	Owen-Smith et al., 2005; Pierce et al., 2012
Late-Winter Palmer Drought Severity Index (PRISM ^b)	4 km (Monthly)	Correlated with tick abundance subsequent winter	Home range value of Palmer Drought Severity Index experienced across all winter home ranges during March and April from previous year	Winter	Juvenile and adult survival	DeGiudice et al., 1997; Samuel, 2007
Snow Water Equivalence (DAYMET ^c)	1 km (Daily)	Winter severity	Cumulative amount of water kg/m ² contained in the snowpack from 1 January to 31 May on winter home ranges from previous or current year	Winter	All ^d	Peterson and Allen, 1974; Keech et al., 2000; Parker et al., 2009
Ingesta-free-body fat (nutritional measurements at capture)	Individual (February)	Fecundity, survival	Ultrasound measurement of individual body fat levels for a percent-body fat estimate	None (nutritional measure)	All	Cook et al., 2004; Cook et al., 2013

^aModerate resolution imaging spectroradiometer, normalized difference vegetation index. ^bParameter-elevation regressions on independent slopes model. ^cDaily surface weather and climatological summaries.

^dCumulative monthly values from 1 January to 31 May were used to assess overwinter survival of adults.

(June), and overwinter (January–May) survival of adults. We used the PDSI value from the current year to evaluate survival of (July–February) juveniles. We evaluated the influence of late winter (March and April) drought on the survival of juveniles and adults. We did not evaluate the influence of late winter drought on pregnancy, because mating season for moose typically occurs during the early stages of tick loading, and it would be unlikely that an effect would be detected.

To relate the energetic costs of snow on moose vital rates, we used raster data of Snow Water Equivalent (SWE) from Daymet (Thornton et al., 1997). Although moose are well adapted to deep snow (Coady, 1974), harsh winters reduce mobility and increase energetic demands, negatively influencing survival of juveniles (Peterson, 1977; Keech et al., 2011) and adults (Peterson and Allen, 1974). Additionally, nutritional costs of moving through snow on maternal condition while calves are *in utero* could produce less viable neonates (Schwartz, 1998, pp. 141–171). We calculated cumulative values of daily SWE from moose winter ranges annually from 1 January to 31 May to evaluate the influence of snow on parturition, survival of neonates (SWE experienced while *in utero*), and survival of juveniles (*in utero*, and first winter as a juvenile). We calculated cumulative values of monthly SWE from January to May to evaluate the influence of winter severity on survival of adults.

Endogenous fat reserves are critical for survival and fecundity of ungulates (Cook et al., 2004, 2010; Monteith et al., 2014). We measured and evaluated the influence of IFBfat in addition to the aforementioned environmental covariates on pregnancy, parturition, and survival of juveniles and adults.

Ungulate populations that have senesced in age can confound estimates of survival and fecundity (Festa-Bianchet et al., 2003; Bonenfant et al., 2009). We included age as a covariate for all vital rate analyses, and determined whether a quadratic transformation was appropriate using residual diagnostics.

Statistical Analyses

Vital Rate Models

For probability of pregnancy, parturition, and survival of neonates and juveniles, we fit binomial GLMMs with the logit link function and individual animal as a random intercept. To assess monthly survival of adults, we used the Andersen–Gill formulation (Andersen and Gill, 1982) of Cox proportional hazards CPHs models (Cox, 1972). We used an information-theoretic approach to model selection by assessing all possible combinations of covariates (Doherty et al., 2012) selected *a priori* based on Akaike information criteria adjusted for small sample size (AIC_c), ΔAIC_c , and Akaike weights (w_i) (Burnham and Anderson, 2002). If Pearson correlation coefficients between each pair of covariates were $> |0.5|$, we assessed univariate models between each pair and selected the covariate with the minimum AIC to use for model selection. Consequently, all variables that entered AIC_c model selection had Pearson correlation coefficients $< |0.5|$. To improve convergence of the GLMMs and CPHs, we scaled all covariate values so that their mean was zero and their standard deviation was one. For all analyses, we omitted mortalities attributed to

anthropogenic causes (e.g., harvest, poaching; $n = 4$). During model selection, we followed Babyak (2004), using approximately 10 events (i.e., failures or mortalities) per covariate (hereafter, “EPC”) that entered a model. Model fit was assessed using area under the receiver operating characteristic curve (AUC; Hanley and McNeil, 1982).

We deemed covariates important if their 95% confidence intervals did not overlap zero within ΔAIC_c four of the top model.

In modeling monthly survival of adults, we assessed winter (January–May) and summer (June–December) separately, because we expected environmental covariates to affect moose survival differently with regard to season (Gaillard et al., 2000). We used the Andersen–Gill formulation of CPHs with each month representing a time interval (i.e., the counting process; Therneau and Grambsch, 2000), allowing for left-staggered entry (i.e., newly captured moose added to sample size) and right-censoring of adults (e.g., if the VHF transmitter failed, collar dropped from the individual, or the moose emigrated from study area). Coefficient estimates from CPHs specify the mortality hazard, where positive values increase mortality hazard (i.e., negatively relate to survival). Adults were allowed to contribute to the risk sample every year they were monitored; therefore, we used a robust “sandwich” estimator to account for correlated observations within individuals (Therneau and Grambsch, 2000). We chose a recurrent time-scale for seasonal CPHs based off of the biological year (i.e., 1 June–31 May), because adult mortality occurred largely during late winter (i.e., a strong seasonal hazard; Fieberg and DelGiudice, 2009). We allowed monthly SWE values to cumulatively increase over winter (e.g., a time-varying covariate; Therneau and Grambsch, 2000). Model fit of CPHs was assessed using concordance (Harrell, 2015), which is analogous to AUC. Diagnostic tests were performed on all CPH models to evaluate the proportional hazards assumption using Schoenfeld residuals (Therneau and Grambsch, 2000), but were not reported unless significant violations were detected (i.e., $P < 0.05$).

Life Stage Simulation Analysis

To assess the influence of interannual variation in vital rates on λ , we conducted LSA. We followed Morris and Doak (2002), and estimated vital rates separately by biological year and then derived a beta distribution of 10,000 estimates for each vital rate by randomly sampling the mean and variance estimated from each year. To derive the 10,000 estimates of λ from beta distributions, we used a 3×3 stage-structured, post-birth, female-based matrix model (Caswell, 2001) consisting of the following form:

$$\lambda = \begin{bmatrix} 0 & 0 & (S_a F T) \\ S_j & 0 & 0 \\ 0 & S_y & S_a \end{bmatrix} \quad (1)$$

The first stage calculates reproduction rates by multiplying adult survival (S_a) by fecundity ($F = \text{pregnancy rate} \times \text{parturition rate}$), and probability of twinning [$T = (1-t) + (2t)$].

The second stage is juvenile survival (S_j), calculated as the neonatal (June to July) survival rate $\times$ juvenile survival (July to

Feb) rate. Our surveys of juvenile survival were conducted during February; therefore, to correct for missed late-winter mortality of calves (Gaillard et al., 1998), we assumed that an additional 10% of mortality would occur between February and May 31. Effectively, this additional 10% did not change the variability of the vital rate in the LSA and only influenced estimates of λ . The third stage consists of yearling survival (S_y) and adult survival (S_a). Rates of yearling survival in the Sublette herd were unavailable, so we calculated a mean estimate from a range of yearling survival rates for moose reported in Kunkel and Pletscher (1999) and McLaren et al. (2000). On average, yearling survival was 6.43% lower than adult survival, thus our estimate was calculated as $S_y = S_a - 6.43\%$. We assumed no reproduction of yearlings in our model. The resulting beta distributions of each vital rate were randomly sampled for vital rate estimates to produce 10,000 matrix replicates, estimates of λ , elasticity and sensitivity. Elasticity analysis estimates the effect of a proportional change in the vital rates on population growth rate, whereas sensitivity analysis estimates the impact of an absolute change in vital rates on λ . We then regressed λ on each vital rate, providing an estimate of the proportion of variation in λ explained by each vital rate (r^2). We were unable to account for sampling variance due to the limited time-frame of the study.

Covariate LSA

After assessing vital rate variability and its influence on λ , we evaluated the influence of the covariates on λ through their relative effect on vital rates. To estimate the strength and variability of covariates on vital rates, we extended the LSAs to incorporate the effects of covariates on the vital rates. We estimated rates of pregnancy, parturition, neonatal, juvenile, and overwinter adult survival as a function of the coefficient estimates of the covariates for 10,000 iterations, and populated the same 3×3 stage-structured, female-based, post-birth matrix models used in LSAs with these predicted values (hereafter, “covariate LSA”). For the covariate LSA, we used estimates of adult survival from the overwinter (January–May) model only, because summer models failed to converge due to small sample size of mortalities. We accounted for total annual mortality by subtracting the average rate of summer (June–December) mortality from each predicted estimate of overwinter survival. For vital rate responses that were not predicted by any covariate, we applied the beta distribution from the LSA to account for unexplained variation. Yearling survival was also set as a constant, estimated at 6.43% lower than adult survival ($S_y = S_a - 6.43\%$). Similar to the LSA, we then regressed λ on each covariate, providing an estimate of the amount of variation (r^2) in λ explained by the covariate.

RESULTS

Probabilities of pregnancy, parturition, and overwinter survival of adults were positively and strongly related to IFBFat (Figures 2A–C and Table 3). Drought conditions on summer home ranges from the previous year reduced probability of

parturition (Figure 2D and Table 3). We did not detect important predictors of juvenile survival. Variation in survival of adults was a strong driver of interannual changes in λ ($r^2 = 0.70$; Figures 3A,F). Variation in neonate and juvenile survival, as well as rates of parturition, and pregnancy contributed comparatively little to population growth (Figures 3B–F). Elasticity values for fecundity and survival of juveniles were 0.12, compared with 0.64 for survival of adults. Sensitivity values for fecundity, survival of juveniles, and survival of adults were 0.30, 0.25, and 0.75, respectively. From the time period of the demographic rates measured, the Sublette moose population was increasing ($\lambda = 1.027$). Twinning rates were low, neonatal survival remained stable and high, and overwinter survival of juveniles was stable, but slightly reduced (Table 2). Although age competed with other predictors in top models for probability of parturition and pregnancy, all confidence intervals overlapped zero (Table 3). Age as a sole predictor of adult survival was important (CPH $\beta = 0.73$; 95% CI: 0.41–1.05), but was clearly outweighed by IFBFat (Table 3). Prime-aged (2–8 years) moose accounted for 63% (SE ± 0.10 ; $n = 15$) of total adult mortalities ($n = 24$), and 84% of all individuals were ≤ 8 years in age ($\bar{x} = 6$; range: 2.5–15 years).

The covariate LSA revealed a strong influence of IFBFat on population growth rate, explaining 82% of variation in λ through its combined influence on overwinter survival of adults, parturition, and pregnancy (Figure 4). Only 0.7% of the variation in λ was explained by drought conditions from the previous summer.

DISCUSSION

Our study illustrates how measurements of nutritional condition can better characterize degree of resource limitation in populations of a large herbivore than remotely sensed metrics of environmental variation. Although we did not measure density dependence directly, we suspect that the influence of nutritional condition at least partially reflects a key role of density dependence operating within a population of moose largely free of large predators with low female harvest—a conclusion that would have been impossible to draw based on inference from remotely sensed environmental measures as proxies of resource limitation. Lambda was explained almost entirely by nutritional condition of females, which clearly illustrated that this population was resource limited and likely experiencing density dependence at a level that obscured the effects of environmental variation detected from remotely sensed data. Moreover, response and variability of vital rates to resource limitation for a population of moose at the southern extent of its distribution conflicted with the Eberhardt paradigm of life history for long-lived ungulates. Despite the typical vulnerability of juveniles and robustness of adults to resource limitation (Gaillard and Yoccoz, 2003), nutritional condition of adults strongly influenced survival, yet survival of juveniles had little predictive power. Further, our LSA results depart from typical expectations for the life-history characteristics of ungulates, where variation in juvenile survival typically fluctuates

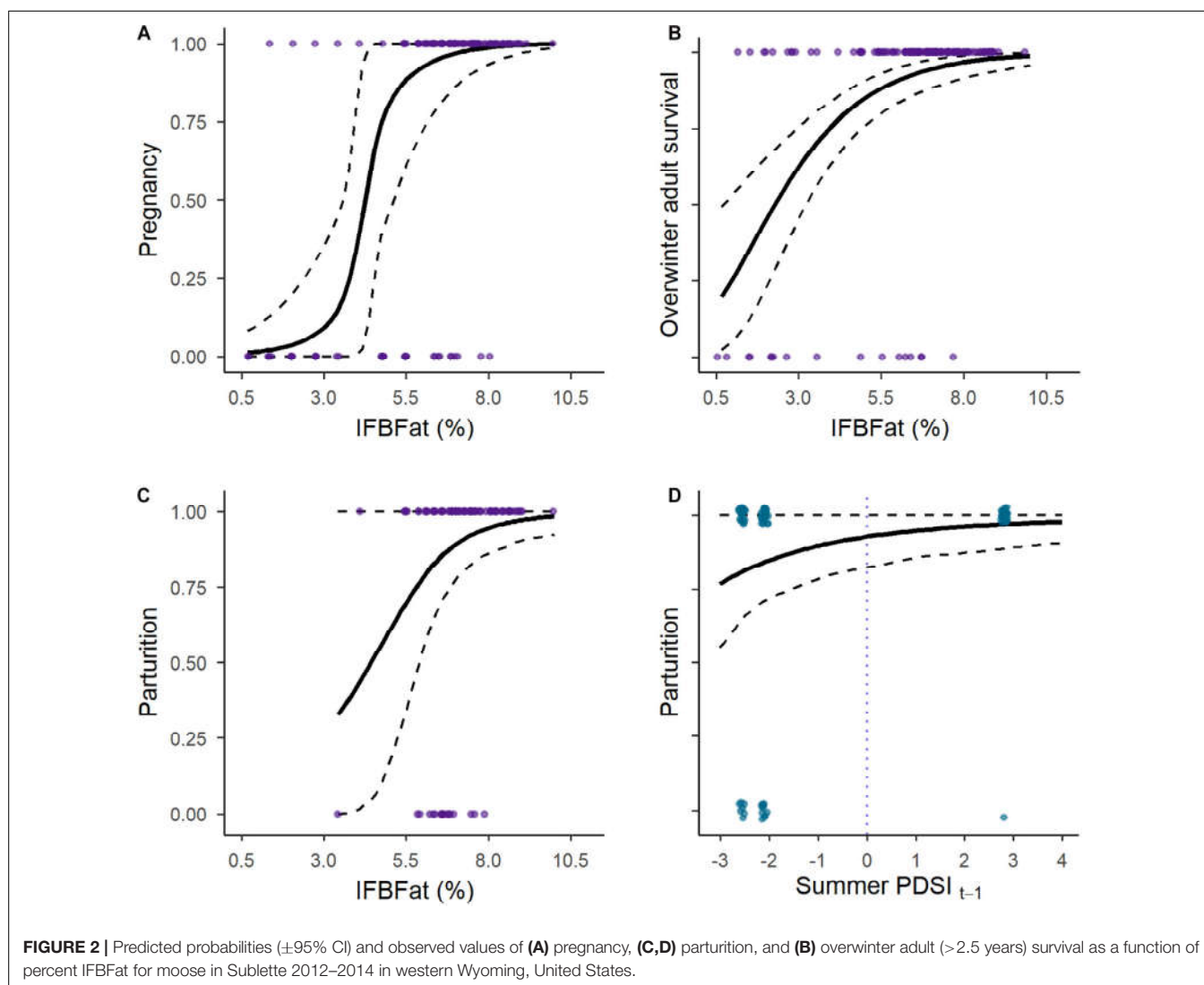


TABLE 2 | Mean estimates (SE) and sample sizes of demographic rates from 2011 to 2014 for Sublette moose in western Wyoming, United States.

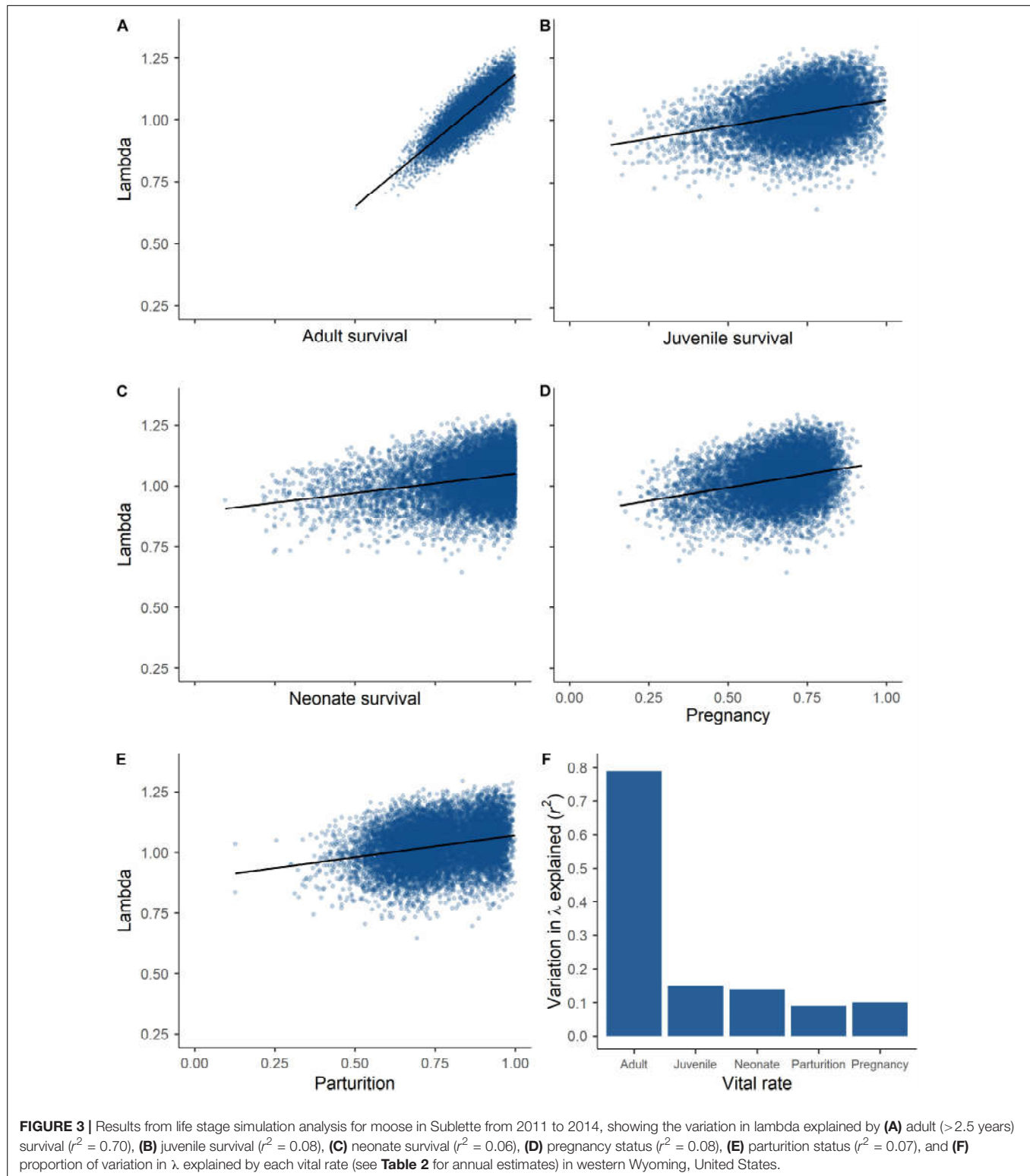
Study	Year	Pregnancy		Parturition		Twinning		Neonate survival		Juvenile survival		Adult survival	
		<i>n</i>	$\bar{x}$ ($\pm$ SE)	<i>n</i>	$\bar{x}$ ($\pm$ SE)	<i>n</i>	$\bar{x}$ ($\pm$ SE)	<i>n</i>	$\bar{x}$ ($\pm$ SE)	<i>n</i>	$\bar{x}$ ($\pm$ SE)	<i>n</i>	$\bar{x}$ ($\pm$ SE)
Sublette	2011	19	0.53 (0.12)	10	0.80 (0.13)	9	0	9	0.71 (0.17)	9	na	23	0.91 (0.06)
	2012	47	0.66 (0.07)	28	0.89 (0.06)	26	0.12 (0.06)	28	0.82 (0.07)	9	0.71 (0.17)	50	0.88 (0.05)
	2013	63	0.76 (0.05)	39	0.69 (0.07)	41	0	30	0.93 (0.05)	25	0.72 (0.09)	70	0.79 (0.05)
	2014	40	0.68 (0.08)	30	0.67 (0.09)	27	0.04 (0.04)	22	0.95 (0.04)	22	0.77 (0.09)	53	0.83 (0.06)
	All years	169	0.69 (0.03)	107	0.75 (0.04)	103	0.04 (0.19)	89	0.89 (0.03)	65	0.74 (0.06)	196	0.84 (0.02)

Rates of pregnancy, parturition, and twinning were calculated as proportions, and survival estimates of neonates, juveniles, and adults are from Kaplan–Meier analysis, with biological year starting 1 June and ending 31 May.

widely and underpins variation in population growth (Gaillard et al., 1998, 2000). In our study, annual changes in rates of population growth were driven mostly by variation in survival of adults, whereas survival of juveniles explained substantially less variation. Our findings serve as a reminder that the influences of top-down and bottom-up forces are context-dependent relative to a population's proximity to K , and that the relative influence of

each factor is obscured or magnified depending on the strength of density dependence.

The detection of resource limitation in Sublette moose through the lens of vital rates more clearly revealed the potential for interacting relationships from top-down and bottom-up forces, and the importance of context dependence in ungulate demography. In our study, signals of drought influencing



population growth rates were barely detected, yet we can logically infer that drought negatively influenced levels of body fat in moose. Warm temperatures during spring and summer reduce forage quality for ungulates (Hamel et al., 2009b),

thereby limiting fat gains during the growing season that are necessary for reproduction and overwinter survival (Monteith et al., 2014). Indeed, in response to drought, Sublette moose experienced reduced rates of parturition, as well as low rates

TABLE 3 | Top models derived from AIC_c model selection with coefficient estimates and 95% confidence intervals (of all terms in model) that did not overlap zero for Sublette moose (*n* = 70 individuals) from 2012 to 2014.

Vital rate	Model	Parameter (β)	AIC _c	Δ	<i>w_i</i>	95% CI	Model fit
Pregnancy	1	Fat ^a (2.39)	111.5	0	0.262	Fat (1.47, 27.15)	0.97
	2	Age ^b (−0.48) + Fat (2.13)	111.8	0.35	0.220		
	3	Fat (2.47) + GrowSeas ^c (−0.33)	112.7	1.23	0.142		
	4	Age (−0.51) + Fat (2.21) + GrowSeas (−0.34)	113	1.51	0.123		
	5	Fat (2.39) + SprLgth ^d (0.19)	113.3	1.82	0.106		
Parturition	1	Fat (0.88) + SmDroughtPY ^e (0.93)	84.7	0	0.585	Fat (0.25, 8.98), SmDroughtPY (0.19, 8.5)	0.89
	2	Fat (0.93)	87.7	3.07	0.126	Fat (0.27, 2.0)	0.89
	3	Age (−0.39) + Fat (0.84)	88.4	3.71	0.091		
	4	Age (−0.45) + SmDroughtPY (0.92)	89.1	4.48	0.062		
	5	SmDroughtPY (0.94)	89.3	4.66	0.057	SmDrought (0.21, 2.08)	
Adult	1	Fat (−1.01)	140.9	0	0.999	Fat (−1.4, −0.61)	0.82
	2	Age (0.73)	154.1	13.17	0.001	Age (0.41, 1.05)	0.73
	3	SmDroughtPY (−0.5426)	163.4	22.46	0		
	4	Intercept	164.7	23.79	0		
	5	GrowSeas (−0.29)	165.6	24.7	0		

^aIFBFat calculated from rump fat measurements during February captures. ^bAge of individual, in years. ^cLength of the growing season, in days. ^dLength of the spring, in days. ^ePalmer Drought Severity Index from the previous summer's home range (positive values are wetter conditions). Model fit of generalized linear mixed-effects models (GLMMs) for probability of pregnancy and parturition was calculated using the area under the receiver operating characteristic curve. Model fit from Cox proportional hazards models (CPH) of adult survival was calculated via concordance. Negative beta coefficients from CPHs decrease the mortality hazard (i.e., positively relate to survival). We did not detect important covariates for juvenile survival of Sublette moose.

of twinning (Table 2), a symptom of resource limitation in moose (Boertje et al., 2007). A lack of predators combined with strong resource limitation revealed strong relationships between nutritional condition and vital rates. Measurements of nutritional condition thus simultaneously integrate disparate influences on resource limitation, and as a result, they are more informative than remotely sensed covariates in linking bottom-up effects to ungulate demography.

As prey populations approach *K* and competition for food increases, nutritional condition declines, and animals are consequently more sensitive to environmental variation (Aanes et al., 2000; Kie et al., 2003; Monteith et al., 2015), parasites, and disease. Effects of remotely sensed data on vital rate variability of moose in Sublette were subtle, and could have led to a conclusion of weak resource limitation in an area that was almost devoid of large predators. Nevertheless, measures of nutritional condition explained most of the variation in survival of adults, pregnancy, and parturition, leading to a cumulative explanation for 82% of the variation in λ (Figure 4). The striking connections between nutritional condition, and both individual and population performance were unequivocally indicative of a population regulated strongly by resource limitation and, presumably, density dependence which apparently obscured signals of environmental influence from remotely sensed data. In other words, measures of nutritional condition were a more precise representation of resource limitation when compared with remotely sensed proxies such as NDVI and drought, even when those proxies were explicitly measured in a spatial and temporal context. There is increasing support for using remotely sensed data to detect resource limitation in ungulates (Tveraa et al., 2007; Pettoirelli et al., 2011; Bastille-Rousseau et al.,

2015; Monteith et al., 2015); however, our work provides evidence that these relationships can go undetected if integrative measures of resource limitation (e.g., nutritional condition) are not considered. Short-term studies such as ours are common for implementing management actions, and remotely sensed data might not be sufficient for detecting strong resource limitation. Similarly, Stewart et al. (2005) noted that signatures of resource limitation can be dampened in populations experiencing heightened density dependence. In Sublette moose, relationships between vital rates and environmental covariates were relatively weak compared with IFBFat, highlighting the potential for Type II errors in relating only remotely sensed covariates to ungulate demography. We recommend that future studies examining the effects of bottom-up forcing on ungulates carefully consider multiple pathways before declaring the absence of either pressure, and we also agree with others (Cook et al., 2004; Monteith et al., 2014; Stephenson et al., 2020) that nutritional condition is an effective indicator of resource limitation because it integrates both environmental variation and density dependence.

Our results depart from the life-history paradigm proposed by Eberhardt (2002), and deviate from life-history characteristics typical of ungulates (Gaillard et al., 2000). Juveniles are typically more sensitive to environmental variation than adults and are commonly the first stage class to experience declines as populations approach *K* (Eberhardt, 2002), ultimately driving interannual variability in λ (Gaillard et al., 2000). Nevertheless, adult survival was highly variable and explained the majority of variation in λ for our study population. Variable survival of adults is considered rare for ungulates, although exceptions have been noted in declining populations exposed to predators, including bighorn sheep (Johnson et al., 2010), mountain caribou

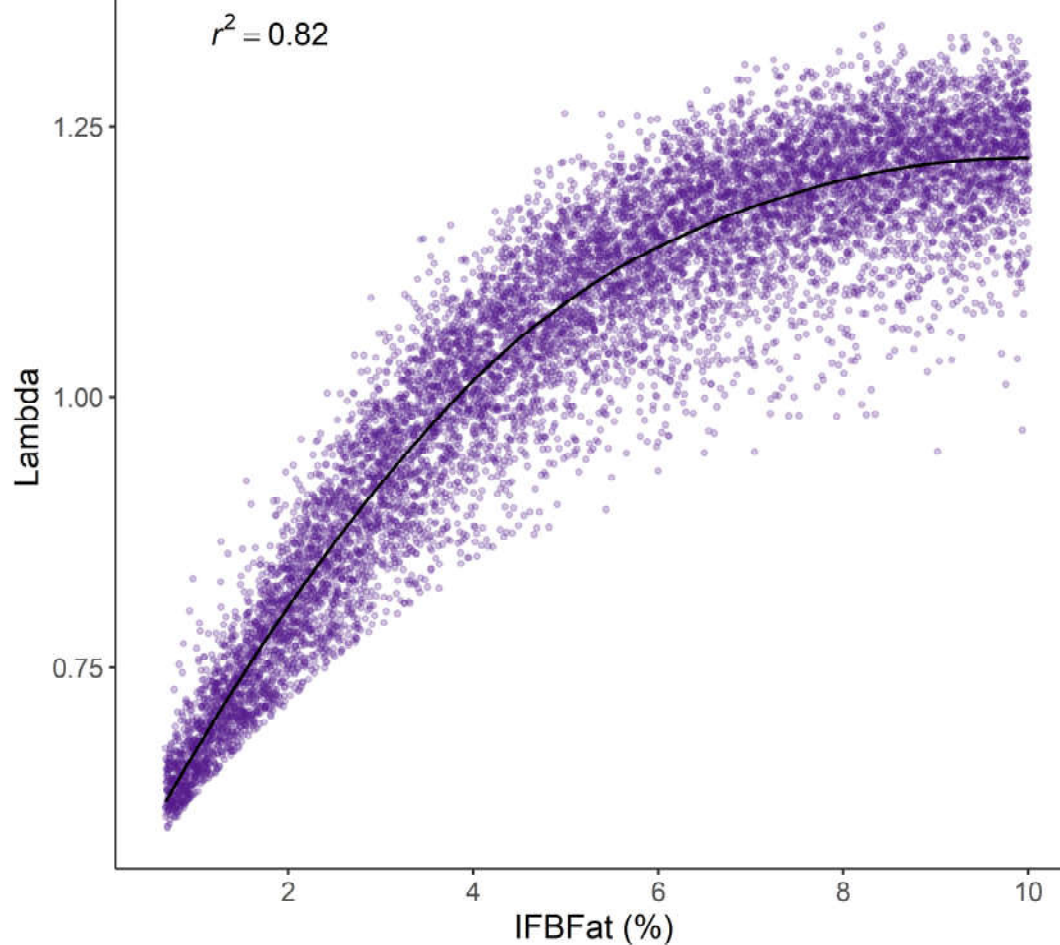


FIGURE 4 | The variation in λ (fitted with smoothing loess) explained by the cumulative effects of percent IFBat on pregnancy, parturition, and overwinter survival of adult (>2.5 years) female moose in Sublette 2012–2014, western Wyoming, United States.

(*Rangifer tarandus*) (Hervieux et al., 2013), and tropical ungulates (Owen-Smith and Mason, 2005). Notably, variable survival of adults in Sublette in response to suppressed nutritional condition suggests that the population surpassed K , despite an estimated stable-to-increasing population growth rate. Moreover, survival of neonates and juveniles was not sensitive to nutritional condition. As intraspecific competition for resources increases, the detection of reproductive tradeoffs (lower juvenile survival) relative to environmental variation may be confounded by factors difficult to measure, such as experience in rearing young (Hamel et al., 2009a), heterogeneity of genotypic quality (van Noordwijk and de Jong, 1986), and immune function (Downs et al., 2015; Cheynel et al., 2017). There are multiple factors that can influence vital rates, and as such, their expected patterns (Gaillard et al., 2000) may be more context dependent than previously appreciated.

Our findings provide three important contributions that can help advance our understanding of resource limitation in ungulate populations. First, in our LSA, variation in adult survival strongly affected lambda, which is notable since

adult survival is typically robust to environmental variation. Second, measurements of nutritional condition proved to be considerably more informative for detecting resource limitation compared with multiple proxies of environmental variation via remote sensing, indicating that seemingly intuitive relationships between proxies of resource limitation and vital rates can be obscured in ungulate populations experiencing strong density dependence. Finally, we documented low variability in survival of young that had little influence on λ , in contrast to expectations (Gaillard et al., 2000). Our results indicate that the life-history paradigm for long-lived ungulates may be more variable across taxa than originally proposed, and moose at the southern extent of their range may uniquely deviate from theoretical expectations.

DATA AVAILABILITY STATEMENT

The datasets generated for this study are available on request to the corresponding author.

ETHICS STATEMENT

The study was reviewed and approved by the University of Wyoming Institutional Animal Care and Use Committee.

AUTHOR CONTRIBUTIONS

BO, MK, KM, JG, and JM conceived the ideas and designed methodology. BO analyzed the data and led the writing of the manuscript. All authors collected the data, critically contributed to the drafts, and gave final approval for publication.

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Conflict of Interest: The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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